# A CONSISTENCY THEOREM FOR REGULAR CONDITIONAL DISTRIBUTIONS

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ABSTRACT. Let  $(\Omega, \mathcal{B})$  be a measurable space,  $\mathcal{A}_n \subset \mathcal{B}$  a sub- $\sigma$ -field and  $\mu_n$  a random probability measure on  $(\Omega, \mathcal{B}), n \geq 1$ . In various frameworks, one looks for a probability P on B such that  $\mu_n$  is a regular conditional distribution for P given  $A_n$  for all n. Conditions for such a P to exist are given. The conditions are quite simple when  $(\Omega, \mathcal{B})$  is a compact Hausdorff space equipped with the Borel or the Baire  $\sigma$ -field (as well as under similar assumptions). Applications to Gibbs measures and Bayesian statistics are given as well.

## 1. The problem

Let  $(\Omega, \mathcal{B})$  be a measurable space and  $\mathbb{P}$  the collection of all probability measures on B. For  $B \in \mathcal{B}$  and any map  $\mu : \Omega \to \mathbb{P}$ , we let  $\mu(B)$  denote the function on  $\Omega$ given by  $\omega \mapsto \mu(\omega)(B)$ . We also let  $\sigma(\mu) = \sigma\{\mu(B) : B \in \mathcal{B}\}.$ 

Let  $P \in \mathbb{P}$  and  $\mathcal{A} \subset \mathcal{B}$  a sub- $\sigma$ -field. A regular conditional distribution (r.c.d.), for P given A, is a map  $\mu : \Omega \to \mathbb{P}$  such that  $\mu(B)$  is a version of  $E_P(I_B \mid A)$  for all  $B \in \mathcal{B}$ . For a r.c.d. to exist, it suffices that P is perfect and B countably generated.

This note originates from the following question. Given a sub- $\sigma$ -field  $\mathcal{A} \subset \mathcal{B}$  and a map  $\mu : \Omega \to \mathbb{P}$  such that  $\sigma(\mu) \subset \mathcal{A}$ , under what conditions is there  $P \in \mathbb{P}$  such that  $\mu$  is a r.c.d. for P given  $\mathcal{A}$ ? Such a question is easily answered. Once stated, however, it grows quickly into the following new question. Suppose we are given  $\{\mathcal{A}_n, \mu_n : n \in I\}$ , where  $\mathcal{A}_n \subset \mathcal{B}$  is a sub- $\sigma$ -field,

$$
\mu_n : \Omega \to \mathbb{P}
$$
 is a map such that  $\sigma(\mu_n) \subset \mathcal{A}_n$ ,

and  $I = \{1, 2, ...\}$  or  $I = \{1, ..., k\}$  for some  $k \geq 1$ . Under what conditions is there  $P \in \mathbb{P}$  such that  $\mu_n$  is a r.c.d. for P given  $\mathcal{A}_n$  for all  $n \in I$ ? If such a P exists, the  $\mu_n$  are said to be *consistent*.

We aim to give conditions for the  $\mu_n$  to be consistent; see Theorems 6 and 7. Throughout, M denotes the (possibly empty) set

 $\mathbb{M} = \{P \in \mathbb{P} : \mu_n \text{ is a r.c.d. for } P \text{ given } \mathcal{A}_n \text{ for all } n \in I\}.$ 

# 2. MOTIVATIONS

Reasonable conditions for consistency, if available, are of potential interest.

As a first (heuristic) example suppose that, for each  $n \in I$ , expert n declares his/her opinions on a certain phenomenon conditionally on his/her information  $\mathcal{A}_n$ . This produces a collection of random probability measures  $\mu_n : \Omega \to \mathbb{P}$  such that  $\sigma(\mu_n) \subset \mathcal{A}_n$ . In this framework, most literature focus on how to summarize the

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experts' opinions  $\mu_n$ . But a preliminary (and not escapable) question is whether the  $\mu_n$  are consistent. If they are not, some of the  $\mu_n$  should be discarded.

Let us turn now to more formal examples.

2.1. Gibbs measures. The consistency problem of Section 1 is classical in statistical mechanics. The following is from [3]. Let  $X = (X_i : i \in S)$  be a process, with state space  $(E, \mathcal{E})$ , indexed by the countable set S. In order to assess the probability distribution of  $X$ , to be called a Gibbs measure, we proceed as follows.

Let  $\Omega = E^S$  and  $\mathcal{B} = \sigma(X_i : i \in S)$ , where  $X_i : \Omega \to E$  is the *i*-th canonical projection. For each finite  $\Lambda \subset S$ , we assign a random probability measure  $\gamma_{\Lambda}$  on  $(E^{\Lambda}, \mathcal{E}^{\Lambda})$ , measurable with respect to  $\sigma(X_i : i \notin \Lambda)$ . Precisely,  $\gamma_{\Lambda}(\omega)$  is a probability measure on  $(E^{\Lambda}, \mathcal{E}^{\Lambda})$  for each  $\omega \in \Omega$  and  $\sigma(\gamma_{\Lambda}) \subset \sigma(X_i : i \notin \Lambda)$ . Here,  $\gamma_{\Lambda}$  is regarded as the conditional distribution of  $(X_i : i \in \Lambda)$  given  $(X_i : i \notin \Lambda)$ . Indeed, for finite  $\Lambda$ , the Gibbsian formalism of equilibrium statistical mechanics provides a simple and reasonable scheme for selecting  $\gamma_{\Lambda}$ . But of course a consistency problem arises for the collection  $\{\gamma_{\Lambda} : \Lambda \text{ finite}\}.$ 

Precisely, fix any enumeration  $\Lambda_1, \Lambda_2, \ldots$  of the finite parts of S and define

$$
Z_n = (X_i : i \notin \Lambda_n), \quad \mathcal{A}_n = \sigma(Z_n), \quad \mu_n(\omega) = \gamma_{\Lambda_n}(\omega) \times \delta_{Z_n(\omega)}.
$$

Then, assessing  $\{\gamma_{\Lambda} : \Lambda \text{ finite}\}\$  makes sense if and only if the  $\mu_n$  are consistent.

2.2. **Bayesian inference.** Loosely speaking, given two events  $A$  and  $B$ , to assess  $Prob(A | B)$  is often easier than to evaluate  $Prob(A)$ . This vague remark can be useful in Bayesian statistics.

Let  $(\mathcal{X}, \mathcal{F})$  and  $(\Theta, \mathcal{G})$  be measurable spaces and  $\mathcal{P} = \{P_{\theta} : \theta \in \Theta\}$  a (measurable) collection of probabilities on  $\mathcal F$ . Fix a prior probability  $\pi$  on  $\mathcal G$  and define  $m(F) = \int P_{\theta}(F) \pi(d\theta)$  for  $F \in \mathcal{F}$ . A posterior for  $\mathcal P$  and  $\pi$  is a (measurable) collection  $\mathcal{Q} = \{Q_x : x \in \mathcal{X}\}\$  of probabilities on  $\mathcal{G}$  satisfying

$$
\int_G P_\theta(F) \,\pi(d\theta) = \int_F Q_x(G) \, m(dx) \quad \text{for all } F \in \mathcal{F} \text{ and } G \in \mathcal{G}.
$$

The standard Bayes procedure is to select a prior  $\pi$  and to calculate (or to approximate) the posterior  $\mathcal{Q}$ . To assess  $\pi$  is often very arduous. Sometimes, it is more convenient to avoid the explicit choice of  $\pi$ , and to assign directly a collection  $\mathcal{Q} = \{Q_x : x \in \mathcal{X}\}\$  of probabilities on G. Here,  $Q_x$  describes our opinions on  $\theta$ when  $x$  is observed. In these cases, the standard Bayes scheme is reverted. One first selects some collection  $Q$  of probabilities on  $G$  and then asks whether  $Q$  is consistent with  $P$ , in the sense that  $Q$  is the posterior of  $P$  and some prior  $\pi$ .

Examples of this "reverted" Bayes procedure are not unusual. Suppose  $\mathcal Q$  is the formal posterior of an improper prior, or else it is obtained by some empirical Bayes method. Then,  $\mathcal{Q}$  is assessed without explicitly selecting a (proper) prior  $\pi$ . Such Q may look reasonable or not (there are indeed different opinions). But again, a preliminary (and not escapable) question is whether  $Q$  is consistent with  $P$ .

The reverted Bayes procedure agrees with the subjective view of probability. In fact, it has been developed in a coherence framework; see [1], [4], [6], [7] and references therein. However, in a coherence framework,  $P$  and  $Q$  are requested to be consistent under some finitely additive prior.

To investigate the reverted Bayes procedure, without using finitely additive probabilities but relying on standard (Kolmogorovian) probability theory, we need exactly the notion of consistency of Section 1.

### 3. RESULTS

3.1. Preliminaries. We begin with a couple of lemmas and a corollary.

**Lemma 1.** Let  $\mathcal{A} = \sigma(\cup_{n \in I} A_n)$ . The  $\mu_n$  are consistent if and only if

(1) 
$$
Q(\mu_n(A) = I_A) = 1 \quad \text{whenever } n \in I \text{ and } A \in \mathcal{A}_n,
$$

(2) 
$$
E_Q\{\mu_n(B)\} = E_Q\{\mu_1(B)\} \text{ whenever } n \in I \text{ and } B \in \mathcal{B},
$$

for some probability measure Q on A. In particular, if  $Q$  meets conditions  $(1)-(2)$ and  $P(\cdot) = E_Q\{\mu_1(\cdot)\}\$ , then  $P \in \mathbb{M}$  and  $P = Q$  on  $\cup_{n \in I} A_n$ . Moreover, if each  $A_n$ is countably generated, condition (1) can be written as  $Q(\Omega_0) = 1$ , where

$$
\Omega_0 = \{ \mu_n(A) = I_A \text{ for all } n \in I \text{ and } A \in \mathcal{A}_n \}.
$$

*Proof.* If the  $\mu_n$  are consistent, it suffices to let  $Q = P | A$  for some  $P \in M$ . Conversely, suppose conditions (1)-(2) hold for some Q. Define  $P(B) = E_{Q} \{\mu_1(B)\}\$ for all  $B \in \mathcal{B}$ . Fix  $n \in I$ ,  $A \in \mathcal{A}_n$  and  $B \in \mathcal{B}$ . Then,  $\mu_n(A \cap B) = I_A \mu_n(B)$  on the set  $\{\mu_n(A) = I_A\}$ . Thus, conditions (1)-(2) yield

$$
P(A \cap B) = E_Q \{ \mu_1(A \cap B) \} = E_Q \{ \mu_n(A \cap B) I_{\{\mu_n(A) = I_A\}} \} = E_Q \{ I_A \mu_n(B) \}.
$$

For  $B = \Omega$ , the above equation reduces to  $P(A) = Q(A)$ . Hence,  $P = Q$  on  $\mathcal{A}_n$ . Since  $\sigma(\mu_n) \subset \mathcal{A}_n$ , it follows that  $E_P\{I_A \mu_n(B)\} = E_Q\{I_A \mu_n(B)\} = P(A \cap B)$ . This proves that  $P \in \mathbb{M}$ . Finally, suppose the  $\mathcal{A}_n$  countably generated and take countable fields  $\mathcal{U}_n$  such that  $\mathcal{A}_n = \sigma(\mathcal{U}_n)$ . Since  $\Omega_0 = \bigcap_{n \in I} \bigcap_{A \in \mathcal{U}_n} {\{\mu_n(A) = I_A\}}$ , then  $\Omega_0 \in \mathcal{A}$  and condition (1) amounts to  $Q(\Omega_0) = 1$ .

In a sense, up to replacing  $\Omega$  with  $\Omega_0$ , condition (1) can be assumed to be true whenever the  $A_n$  are countably generated. In this case, in fact, the  $\mu_n$  are certainly not consistent if  $\Omega_0 = \emptyset$ . Otherwise, if  $\Omega_0 \neq \emptyset$ , they are consistent if and only if there is a probability Q on the trace  $\sigma$ -field  $\mathcal{A} \cap \Omega_0$  satisfying condition (2).

Among other things, Lemma 1 answers our initial question, raised in Section 1. Suppose in fact  $I = \{1\}$ . Since (2) is trivially true,  $\mathbb{M} \neq \emptyset$  if and only if condition (1) holds. In turn, condition (1) is automatically true if  $\Omega_0 \neq \emptyset$  (just let  $Q = \delta_{\omega}$ for some  $\omega \in \Omega_0$ , so that Lemma 1 implies  $\mu_1(\omega) \in \mathbb{M}$ ). Furthermore,  $\Omega_0 \neq \emptyset$  is equivalent to (1) if  $A_1$  is countably generated (but not in general).

From now on, whether or not the  $A_n$  are countably generated, it is assumed  $\Omega_0 = \Omega$  or equivalently

(3) 
$$
\mu_n(\omega)(A) = I_A(\omega) \text{ for all } \omega \in \Omega, n \in I \text{ and } A \in \mathcal{A}_n.
$$

**Lemma 2.** Suppose condition (3) holds. The  $\mu_n$  are consistent if and only if

(4) 
$$
\int \mu_j(\omega)(B) Q_j(d\omega) = \int \int \mu_n(x)(B) \mu_j(\omega)(dx) Q_j(d\omega)
$$

for some  $j \in I$ , some probability  $Q_j$  on  $A_j$ , and all  $n \in I$  and  $B \in \mathcal{B}$ .

*Proof.* Suppose the  $\mu_n$  are consistent. Fix  $P \in \mathbb{M}$  and  $j \in I$  and let  $Q_j = P | \mathcal{A}_j$ . Since  $P(dx) = \mu_i(\omega)(dx) Q_i(d\omega)$ , then

$$
\int \int \mu_n(x)(B) \mu_j(\omega)(dx) Q_j(d\omega) = \int \mu_n(x)(B) P(dx) = P(B) = \int \mu_j(\omega)(B) Q_j(d\omega)
$$

for all  $n \in I$  and  $B \in \mathcal{B}$ . Conversely, suppose condition (4) holds for some  $j \in I$ and probability  $Q_j$  on  $A_j$ . Define  $P(B) = E_{Q_j} \{\mu_j(B)\}\$  for all  $B \in \mathcal{B}$ . By (3),

$$
P(A) = E_{Q_j} \{ \mu_j(A) \} = E_{Q_j} \{ I_A \} = Q_j(A) \text{ for all } A \in \mathcal{A}_j.
$$

Let  $n \in I$  and  $B \in \mathcal{B}$ . Since  $P = Q_i$  on  $\mathcal{A}_i$ , condition (4) yields

$$
E_P\{\mu_n(B)\} = \int \mu_n(x)(B) P(dx) = \int \int \mu_n(x)(B) \mu_j(\omega)(dx) Q_j(d\omega)
$$

$$
= \int \mu_j(\omega)(B) Q_j(d\omega) = \int \mu_j(\omega)(B) P(d\omega) = E_P\{\mu_j(B)\}.
$$

An application of Lemma 1 (with  $Q = P|\mathcal{A}$ ) concludes the proof.

**Corollary 3.** Suppose condition (3) holds. If there is  $\omega_0 \in \Omega$  such that

(5) 
$$
\mu_1(\omega_0)(\cdot) = \int \mu_n(x)(\cdot) \mu_1(\omega_0)(dx) \text{ for all } n \in I,
$$

the  $\mu_n$  are consistent. Moreover, condition (5) is equivalent to consistency of the  $\mu_n$  in case B is countably generated and  $\mathcal{A}_n \supset \mathcal{A}_1$  for all  $n \in I$ .

*Proof.* Just apply Lemma 2 with  $j = 1$  and  $Q_j = \delta_{\omega_0}$ . Next, suppose  $\beta$  countably generated,  $A_n \supset A_1$  for all n, and the  $\mu_n$  are consistent. Fix  $P \in \mathbb{M}$  and define  $A = {\omega : \mu_1(\omega)(\cdot) = \int \mu_n(x)(\cdot) \mu_1(\omega)(dx)}$  for all  $n \in I}$ . Then,

$$
\mu_1(B) = E_P(I_B | A_1) = E_P\{E_P(I_B | A_n) | A_1\} = \int \mu_n(x)(B) \mu_1(\cdot)(dx), \quad P\text{-a.s.},
$$

for fixed  $B \in \mathcal{B}$  and  $n \in I$ . Hence,  $\mathcal{B}$  countably generated yields  $P(A) = 1$ .

 $\Box$ 

When the  $\mu_n$  are consistent, various questions on M arise. A natural one is uniqueness of  $P \in \mathbb{M}$ . Another question is existence (and possibly uniqueness) of  $P \in \mathbb{M}$  such that  $P \sim P_0$ , where  $P_0$  is a given reference measure.

In general, to find non trivial conditions for uniqueness of  $P \in M$  looks very arduous. For instance,  $P(\cdot | A) \in \mathbb{M}$  whenever  $P \in \mathbb{M}$ ,  $A \in \cap_n \mathcal{A}_n$  and  $P(A) > 0$ . However, uniqueness conditions are available in particular cases. One such case is that of Gibbs measures; see Chapter 8 of [3]. Here, incidentally, uniqueness of  $P \in \mathbb{M}$  is crucial as non uniqueness corresponds to phase transitions.

The second question is connected to equivalent martingale measures.

**Proposition 4.** Suppose condition (3) holds and  $\cup_{n\in I} A_n$  is a field. Fix a probability P<sub>0</sub> on  $\mathcal{A} = \sigma(\cup_{n \in I} A_n)$  and let F be the linear space generated by  $\mu_n(B) - \mu_1(B)$ for all  $n \in I$  and  $B \in \mathcal{B}$ . Then, there is  $P \in \mathbb{M}$  such that  $P \sim P_0$  on A if and only if

(6) 
$$
\overline{F - L_{\infty}^{+}} \cap L_{\infty}^{+} = \{0\}
$$

where  $L_{\infty} = L_{\infty}(\Omega, \mathcal{A}, P_0)$  and the closure is in the weak\* topology on  $L_{\infty}$ .

Proof. By a result of Kreps [5], condition (6) is equivalent to existence of a probability  $Q$  on  $A$  such that

(7) 
$$
Q \sim P_0
$$
 and  $E_Q(f) = 0$  for all  $f \in F$ .

Thus, if  $P \in \mathbb{M}$  and  $P \sim P_0$  on A, it suffices to note that  $Q = P | A$  meets condition (7). Conversely, under (6), take Q satisfying (7) and define  $P(\cdot) = E_Q\{\mu_1(\cdot)\}\$ . In view of (3), Lemma 1 implies  $P \in \mathbb{M}$  and  $P = Q$  on  $\cup_{n \in I} A_n$ . Since  $\cup_{n \in I} A_n$  is a field, it follows that  $P = Q \sim P_0$  on A.

As an extreme example, if F is finite dimensional, existence of  $P \in \mathbb{M}$  such that  $P \sim P_0$  on A reduces to the no arbitrage condition

$$
P_0(f > 0) > 0 \iff P_0(f < 0) > 0
$$
 for each  $f \in F$ .

3.2. Main results. Some notation is needed. Let  $f : \Omega \to \mathbb{R}$  be a bounded Bmeasurable function. We write  $P(f) = E_P(f) = \int f dP$  whenever  $P \in \mathbb{P}$ . For any map  $\mu : \Omega \to \mathbb{P}$ , we denote  $\mu(f)$  the function on  $\Omega$  given by

$$
\mu(\omega)(f) = \int f(x) \,\mu(\omega)(dx), \quad \omega \in \Omega.
$$

A B-determining class is a class S of bounded B-measurable functions  $f : \Omega \to \mathbb{R}$ such that, for arbitrary  $P_1, P_2 \in \mathbb{P}$ ,

$$
P_1 = P_2 \iff P_1(f) = P_2(f)
$$
 for all  $f \in \mathcal{S}$ .

If X is a topological space,  $C(X)$  denotes the set of real continuous functions,  $\mathfrak{B}(X)$  the Borel  $\sigma$ -field, and  $\mathfrak{B}_0(X) := \sigma[C(X)]$  the Baire  $\sigma$ -field. Say that X is pseudocompact if each  $f \in C(X)$  is a bounded function. Clearly, a compact space is pseudocompact. Our main tool is the following.

**Lemma 5.** Let L be a vector lattice of real functions on a set X. Assume  $1 \in L$ and, for any function  $f: X \to [0, \infty)$ ,

$$
f \land n \in L
$$
 for all  $n \ge 1 \implies f \in L$ .

Then, each linear positive functional U on L admits the representation  $U(f) = \int f d\nu$ ,  $f \in L$ , for some (unique) measure  $\nu$  on  $\sigma(L)$ . Next, suppose every  $f \in L$  is bounded and fix a linear subspace  $F \subset L$ . If sup  $f \geq 0$  for all  $f \in F$ , there is a probability measure P on  $\sigma(L)$  such that  $E_P(f) = 0$  for all  $f \in F$ .

Proof. The first part of the Lemma is Theorem 8 of [2]. We prove the second part. Suppose  $f_1 + \lambda_1 = f_2 + \lambda_2$  where  $f_1, f_2 \in F$  and  $\lambda_1, \lambda_2 \in \mathbb{R}$ . Since F is a linear space and sup  $f \ge 0$  for all  $f \in F$ , one obtains  $\lambda_1 = \lambda_2$  and  $f_1 = f_2$ . Let G be the linear space generated by F and the constants. Each  $q \in G$  admits the representation  $g = f + \lambda$ , for some unique  $f \in F$  and  $\lambda \in \mathbb{R}$ , and thus one can define  $T(q) = T(f + \lambda) = \lambda$ . Such a T is a linear functional on G satisfying  $T = 0$  on F and  $T(q) \leq \sup g$  for all  $g \in G$ . Since L consists of bounded functions,  $f \mapsto \sup f$ is a real sublinear functional on  $L$ . By Hahn Banach theorem,  $T$  can be extended to a linear functional U on L satisfying  $U(f) \leq \sup f$  for all  $f \in L$ . Finally, as  $U(1) = T(1) = 1$ , the first part of the Lemma yields  $U(f) = E_P(f)$ ,  $f \in L$ , for some probability P on  $\sigma(L)$ .

Note that, if X is a pseudocompact space and  $\mathcal E$  a  $\sigma$ -field on X, one can take  $L = C(X)$  or  $L = \{f \in C(X) : f \text{ is } \mathcal{E}\text{-measurable}\}\$ in Lemma 5. We are now in a position to state our main result.

**Theorem 6.** Suppose condition (3) holds. Fix  $j \in I$  and a B-determining class S. The  $\mu_n$  are consistent provided one of the following conditions (a)-(b) holds.

(a)  $\Omega$  is a pseudocompact space and  $\mathcal{B} \subset \mathfrak{B}_0(\Omega)$ . Further,

$$
\sup_{\omega \in \Omega} h(\omega) \ge 0 \quad and \quad h \in C(\Omega)
$$

for each function  $h : \Omega \to \mathbb{R}$  in the linear space generated by

$$
\{\mu_n(f) - \mu_j(f) : n \in I, f \in \mathcal{S}\}.
$$

(b) There are a pseudocompact space K, a  $\sigma$ -field  $\mathcal{K} \subset \mathfrak{B}_0(K)$  and a surjective map  $\phi : \Omega \to K$ , such that  $\mathcal{A}_j = \phi^{-1}(\mathcal{K})$ . Further,

 $\sup_{\omega \in \Omega} h(\omega) \geq 0$  and h is continuous in the topology induced by  $\phi$ 

for each function  $h : \Omega \to \mathbb{R}$  in the linear space generated by

$$
\Big\{ \int \mu_n(x)(f) \mu_j(\cdot)(dx) - \mu_j(\cdot)(f) : n \in I, f \in S \Big\}.
$$

(The topology induced by  $\phi$  is  $\phi^{-1}(\mathcal{U})$  where  $\mathcal U$  is the topology on K).

*Proof.* We first prove (b). Fix  $h \in H$ , where H is the linear space generated by

$$
\int \mu_n(x)(f) \mu_j(\cdot)(dx) - \mu_j(\cdot)(f) \quad \text{for all } n \in I \text{ and } f \in \mathcal{S}.
$$

Since  $\phi$  is surjective and h measurable with respect to  $\mathcal{A}_j = \phi^{-1}(\mathcal{K})$ , there is a unique function  $k : K \to \mathbb{R}$  such that  $h = k \circ \phi$ . (Just set  $k(a) = h(\omega_a)$  for each  $a \in K$ , where  $\omega_a \in \Omega$  satisfies  $\phi(\omega_a) = a$ . Also  $k \in C(K)$ , due to h is continuous in the topology induced by  $\phi$ , and  $\sup_{a \in K} k(a) = \sup_{\omega \in \Omega} h(\omega) \geq 0$ . Thus,  $F := \{k : h = k \circ \phi \text{ for some } h \in H\}$  is a linear subspace of  $C(K)$  and  $\sup k \geq 0$  for all  $k \in F$ . By Lemma 5, applied with  $X = K$  and  $L = C(K)$ , there is a probability P on  $\sigma(L) = \mathfrak{B}_0(K)$  such that  $E_P(k) = 0$  for all  $k \in F$ . Next, since  $\mathcal{A}_j = \phi^{-1}(\mathcal{K})$  with  $\phi$  surjective, each  $A \in \mathcal{A}_j$  can be written as  $A = {\phi \in B_A}$  for some unique  $B_A \in \mathcal{K} \subset \mathfrak{B}_0(K)$ . Hence, it makes sense to define

$$
Q_j(A) = Q_j(\phi \in B_A) = P(B_A) \text{ for all } A \in \mathcal{A}_j.
$$

Then,  $Q_j$  is a probability on  $\mathcal{A}_j$  and  $Q_j \circ \phi^{-1} = P$  on K. Given  $h \in H$ ,

$$
E_{Q_j}(h) = E_{Q_j}(k \circ \phi) = E_P(k) = 0
$$

where  $k \in F$  and  $h = k \circ \phi$ . In particular,

$$
\int \mu_j(\omega)(f) Q_j(d\omega) = \int \int \mu_n(x)(f) \mu_j(\omega)(dx) Q_j(d\omega) \text{ for all } n \in I \text{ and } f \in \mathcal{S}.
$$

Since S is B-determining, an application of Lemma 2 concludes the proof of  $(b)$ .

Finally, to prove (a), take F the linear space generated by  $\mu_n(f) - \mu_i(f)$  for all  $n \in I$  and  $f \in S$ . Then, (a) follows precisely as (b), by applying Lemma 5 with  $X = \Omega$  and  $L = C(\Omega)$  and by using Lemma 1 instead of Lemma 2.

 $\Box$ 

In real problems, part (a) of Theorem 6 is much more convenient when  $\Omega$  is pseudocompact. For non pseudocompact Ω, however, part (b) may be useful as well; see Examples 9 and 11.

The connections between Theorem 6 and Proposition 4 should also be stressed. In a sense, the latter is the density-counterpart of the former. Apart from technicalities, in both cases, the main issue is existence of a probability measure with null expectation on a suitable linear space of bounded random variables. This is

achieved via Lemma 5, as regards Theorem 6, and by a result of Kreps [5] in case of Proposition 4. Perhaps, Theorem 6 and Proposition 4 could be unified in a single statement, or at least they could be given essentially parallel proofs.

Our last result, suggested by ideas in [3] (see Theorems (4.17) and (4.22)), is tailor-made for Gibbs measures.

**Theorem 7.** Suppose condition (3) holds,  $\Omega$  is a pseudocompact space, and

$$
\mathcal{B} = \sigma(V) \quad where \quad V = \{ f \in C(\Omega) : f \text{ is } \mathcal{B}\text{-measurable} \}.
$$

Suppose also that  $\mu_n(f) \in C(\Omega)$  for all  $n \in I$  and  $f \in V$ , and there is a subsequence  ${n_i}$  satisfying

(8) for each  $n \geq 1$ , there is  $j_0 \geq 1$  such that  $\mathcal{A}_{n_i} \subset \mathcal{A}_n$  for all  $j \geq j_0$ .

The  $\mu_n$  are consistent if there is  $\omega_0 \in \Omega$  such that

$$
\mu_n(\omega_0)(\cdot) = \int \mu_m(x)(\cdot) \mu_n(\omega_0)(dx) \quad \text{whenever } \mathcal{A}_n \subset \mathcal{A}_m.
$$

Moreover, existence of such  $\omega_0$  is equivalent to consistency of the  $\mu_n$  in case  $\beta$  is countably generated.

*Proof.* As  $\Omega$  is pseudocompact, each  $f \in V$  is bounded. Since  $\sigma(V) = \mathcal{B}$  and V is closed under multiplications,  $V$  is a  $\beta$ -determining class. By the latter fact and condition (3), for every  $P \in \mathbb{P}$  one obtains

$$
P \in \mathbb{M} \iff E_P\{\mu_n(f)\} = E_P(f)
$$
 for all  $n \in I$  and  $f \in V$ .

Let  $W = \{f \in V : 0 \leq f \leq 1\}$  and let  $[0,1]^W$  be equipped with the product topology. Every  $P \in \mathbb{P}$  can be regarded as a map  $P : W \to [0, 1]$ . Since  $[0, 1]^W$ is compact,  $\{\mu_{n_j}(\omega_0) : j \geq 1\}$  admits a converging subnet, say  $\{\mu_\alpha(\omega_0) : \alpha \in D\}$ where D is a suitable directed set. Each  $f \in V$  can be written as  $f = a + bg$  for some  $a, b \in \mathbb{R}$  and  $g \in W$ . Accordingly, one can define

$$
U(f) = \lim_{\alpha} \mu_{\alpha}(\omega_0)(f) = a + b \lim_{\alpha} \mu_{\alpha}(\omega_0)(g) \quad \text{for all } f \in V.
$$

Such U is a linear positive functional on V with  $U(1) = 1$ . By Lemma 5, applied with  $X = \Omega$  and  $L = V$ , one obtains  $U(f) = E_{P_0}(f)$  for all  $f \in V$  and some probability  $P_0$  on  $\sigma(V) = \mathcal{B}$ . Fix  $n \in I$  and  $f \in V$ . Since  $\mu_n(f) \in V$  and  $\mathcal{A}_\alpha \subset \mathcal{A}_n$ for large  $\alpha$ , then

$$
E_{P_0}\{\mu_n(f)\} = \lim_{\alpha} \mu_\alpha(\omega_0)\{\mu_n(f)\} = \lim_{\alpha} \mu_\alpha(\omega_0)(f) = U(f) = E_{P_0}(f).
$$

Hence,  $P_0 \in \mathbb{M}$  and the  $\mu_n$  are consistent. Finally, suppose  $\beta$  countably generated and define  $A = {\omega : \mu_n(\omega)(\cdot) = \int \mu_m(x)(\cdot) \mu_n(\omega)(dx)}$  whenever  $A_n \subset A_m$ . Then,  $P(A) = 1$  for each  $P \in \mathbb{M}$ , by the same argument in the proof of Corollary 3.

 $\Box$ 

**Remark 8.** In part (a) of Theorem 6,  $\mathcal{B} \subset \mathfrak{B}_0(\Omega)$  can be replaced by  $\mathcal{B} \subset \mathfrak{B}(\Omega)$ if each probability on  $\mathfrak{B}_0(\Omega)$  can be extended to a probability on  $\mathfrak{B}(\Omega)$ . This is trivially true if  $\Omega$  is metric  $(\mathfrak{B}_0(\Omega) = \mathfrak{B}(\Omega)$  in this case) or if  $\Omega$  is compact and Hausdorff. The same comment holds for part (b) up to replacing  $(\Omega, \mathcal{B})$  with  $(K, \mathcal{K})$ .

### 4. Examples

**Example 9.** Let  $\Omega = \mathbb{R}^n \setminus \{0\}$  and  $\mathcal{B} = \mathfrak{B}(\Omega)$ . Define

$$
\phi(\omega) = \frac{\omega}{\|\omega\|}, \quad K = \{\omega : \|\omega\| = 1\}, \quad \mathcal{A}_1 = \phi^{-1}(\mathfrak{B}(K)), \quad \mathcal{S} = \{f \in C(\Omega) : f \text{ bounded}\}.
$$

Define also  $\lambda(\omega) = \frac{\max_i |\omega_i|}{\|\omega\|}$ , where  $\omega_i$  is the *i*-th coordinate of  $\omega$ , and

$$
\mu_1(\omega)(B) = \lambda(\omega) \int_0^\infty I_B[r\,\phi(\omega)] \, \exp(-\lambda(\omega) \, r) \, dr \quad \text{for all } B \in \mathcal{B}.
$$

Then,  $\mu_1(A) = I_A$  if  $A \in \mathcal{A}_1$ . Also, if  $f \in \mathcal{S}$ , it is not hard to see that  $\mu_1(f)$  is continuous in the topology induced by  $\phi$ . Thus, in principle, given any collection  $\{\mu_n : n \in I, n > 1\}$  of random probability measures, consistency of  $\{\mu_n : n \in I\}$ can be checked through part (b) of Theorem 6. To fix ideas, suppose  $A_2 \subset B$  is a sub- $\sigma$ -field and  $\mu_2 : \Omega \to \mathbb{P}$  any map. Then,  $\mu_1$  and  $\mu_2$  are consistent whenever

$$
\sigma(\mu_2) \subset \mathcal{A}_2, \qquad \mu_2(A) = I_A \text{ for } A \in \mathcal{A}_2, \quad \mu_2(f) \in \mathcal{S} \text{ for } f \in \mathcal{S},
$$
  

$$
\sup_{\omega} \lambda(\omega) \int_0^{\infty} \left\{ \mu_2[r\,\phi(\omega)](f) - f[r\,\phi(\omega)] \right\} \exp(-\lambda(\omega) r) \, dr \ge 0 \quad \text{for } f \in \mathcal{S}.
$$

**Example 10. (Gibbs measures).** Let  $(\Omega, \mathcal{B}) = (E^S, \mathcal{E}^S)$  where  $(E, \mathcal{E})$  is a measurable space and  $S$  a countable set. As in Subsection 2.1, select a collection  $\gamma = \{\gamma_{\Lambda} : \Lambda \subset S, \Lambda \text{ finite}\}\$  where each  $\gamma_{\Lambda}$  is a (suitably measurable) random probability measure on  $(E^{\Lambda}, \mathcal{E}^{\Lambda})$ . Given  $\gamma$ , define  $\mathcal{A}_n$  and  $\mu_n$  as in Subsection 2.1. Then, conditions (3) and (8) are automatically true. Therefore, by Theorem 7, the  $\mu_n$  are consistent provided

- (i) E is a compact space and  $\mathcal{E} = \mathfrak{B}_0(E);$
- (ii)  $\mu_n(f) \in C(\Omega)$  for all  $n \geq 1$  and  $f \in V$ , where  $\Omega$  is given the product topology and  $V = \{f \in C(\Omega) : f \text{ is } \mathcal{B}\text{-measurable}\};$

(iii) There is  $\omega_0 \in \Omega$  such that  $\mu_n(\omega_0)(\cdot) = \int \mu_m(x)(\cdot) \mu_n(\omega_0)(dx)$  if  $\mathcal{A}_n \subset \mathcal{A}_m$ .

However, the present example adds very little to what already known. In fact, by arguments in [3] (see e.g. the Introduction to Chapter 4 and Theorem (4.17)), the  $\mu_n$  are consistent provided E is a compact metric space and  $\gamma$  a quasilocal specification. Now, the quasilocality condition essentially amounts to (ii) and each specification  $\gamma$  satisfies

(iv)  $\mu_n(\omega_0)(\cdot) = \int \mu_m(x)(\cdot) \mu_n(\omega_0)(dx)$  if  $\mathcal{A}_n \subset \mathcal{A}_m$  for all  $\omega_0 \in \Omega$ ; we refer to [3] for details. Thus, the only contributions of this example are that metrizzability of  $E$  can be dropped and (iv) can be weakened into (iii).

Example 11. (Bayesian inference). In the notation of Subsection 2.2, define

$$
\Omega = \mathcal{X} \times \Theta, \quad \mathcal{B} = \mathcal{F} \otimes \mathcal{G}, \quad \mathcal{A}_1 = \phi_1^{-1}(\mathcal{F}), \quad \mathcal{A}_2 = \phi_2^{-1}(\mathcal{G}),
$$

$$
\mu_1(x, \theta) = \delta_x \times Q_x, \quad \mu_2(x, \theta) = P_\theta \times \delta_\theta,
$$

where  $\phi_1(x,\theta) = x$  and  $\phi_2(x,\theta) = \theta$  for all  $(x,\theta) \in \mathcal{X} \times \Theta$ . Condition (3) trivially holds. Hence, consistency of  $Q$  with  $P$  can be checked by part (b) of Theorem 6 if at least one between  $(\mathcal{X}, \mathcal{F})$  and  $(\Theta, \mathcal{G})$  is a pseudocompact space equipped with the Baire  $\sigma$ -field. This fact, however, is basically known; see Corollary 3.1 of [6].

**Example 12.** (Predictive inference). For each  $n \in I := \{1, 2, ...\}$ , a point  $x_n$ is observed in a measurable space  $(\mathcal{X}_n, \mathcal{F}_n)$ . The problem is to make inference on

sup

 $(x_{n+1}, x_{n+2}, \ldots)$ , conditionally on  $(x_1, \ldots, x_n)$ , in a sequential framework. Define  $\Omega = \prod_{i=1}^{\infty} X_i$ ,  $\mathcal{B} = \otimes_{i=1}^{\infty} \mathcal{F}_i$  and  $\mathcal{A}_n = \sigma(X_1, \ldots, X_n)$ , where  $X_i(\omega) = x_i$  for all  $\omega = (x_1, \ldots, x_i, \ldots) \in \Omega$ . Also, for each  $n \geq 1$ , select a measurable collection

$$
\mathcal{P}_n = \{ P_n(\cdot \mid x_1, \dots, x_n) : (x_1, \dots, x_n) \in \mathcal{X}_1 \times \dots \times \mathcal{X}_n \}
$$

of laws on  $\otimes_{i>n} \mathcal{F}_i$ . Measurability means that  $(x_1, \ldots, x_n) \mapsto P_n(B \mid x_1, \ldots, x_n)$  is measurable for fixed  $B \in \otimes_{i>n} \mathcal{F}_i$ . Each  $P_n(\cdot \mid x_1, \ldots, x_n)$  should be regarded as the conditional distribution of  $(X_{n+1}, X_{n+2}, ...)$  given that  $X_1 = x_1, ..., X_n = x_n$ . Note that, even if a parameter space  $(\Theta, \mathcal{G})$  is available, we do not assess any prior on G. To test consistency of  $\{\mathcal{P}_n : n \geq 1\}$ , define

$$
\mu_n(\omega) = \delta_{(x_1,\ldots,x_n)} \times P_n(\cdot \mid x_1,\ldots,x_n) \quad \text{where } \omega = (x_1,\ldots,x_n,\ldots) \in \Omega.
$$

Again, condition (3) is trivially true. Thus, by Corollary 3, the  $\mu_n$  are consistent provided condition (5) holds for some  $\omega_0 \in \Omega$ . If the  $\mathcal{F}_n$  are countably generated, existence of  $\omega_0$  satisfying (5) is necessary for consistency as well.

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