

# EXCHANGEABLE SEQUENCES DRIVEN BY AN ABSOLUTELY CONTINUOUS RANDOM MEASURE

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ABSTRACT. Let  $S$  be a Polish space and  $(X_n : n \geq 1)$  an exchangeable sequence of  $S$ -valued random variables. Let  $\alpha_n(\cdot) = P(X_{n+1} \in \cdot \mid X_1, \dots, X_n)$  be the predictive measure and  $\alpha$  a random probability measure on  $S$  such that  $\alpha_n \xrightarrow{\text{weak}} \alpha$  a.s.. Two (related) problems are addressed. One is to give conditions for  $\alpha \ll \lambda$  a.s., where  $\lambda$  is a (non random)  $\sigma$ -finite Borel measure on  $S$ . Such conditions should concern the finite dimensional distributions  $\mathcal{L}(X_1, \dots, X_n)$ ,  $n \geq 1$ , only. The other problem is to investigate whether  $\|\alpha_n - \alpha\| \xrightarrow{\text{a.s.}} 0$ , where  $\|\cdot\|$  is total variation norm. Various results are obtained. Some of them do not require exchangeability, but hold under the weaker assumption that  $(X_n)$  is conditionally identically distributed, in the sense of [2].

## 1. TWO RELATED PROBLEMS

Throughout,  $S$  is a Polish space and

$$X = (X_1, X_2, \dots)$$

a sequence of  $S$ -valued random variables on the probability space  $(\Omega, \mathcal{A}, P)$ . We let  $\mathcal{B}$  denote the Borel  $\sigma$ -field on  $S$  and  $\mathbb{S}$  the set of probability measures on  $\mathcal{B}$ . A random probability measure on  $S$  is a map  $\alpha : \Omega \rightarrow \mathbb{S}$  such that  $\sigma(\alpha) \subset \mathcal{A}$ , where  $\sigma(\alpha)$  is the  $\sigma$ -field on  $\Omega$  generated by  $\omega \mapsto \alpha(\omega)(B)$  for all  $B \in \mathcal{B}$ .

For each  $n \geq 1$ , let  $\alpha_n$  be the  $n$ -th *predictive measure*. Thus,  $\alpha_n$  is a random probability measure on  $S$  and  $\alpha_n(\cdot)(B)$  is a version of  $P(X_{n+1} \in B \mid X_1, \dots, X_n)$  for all  $B \in \mathcal{B}$ . Define also  $\alpha_0(\cdot) = P(X_1 \in \cdot)$ .

If  $X$  is *exchangeable*, as assumed in this section, there is a random probability measure  $\alpha$  on  $S$  such that

$$\alpha_n(\omega) \xrightarrow{\text{weak}} \alpha(\omega) \quad \text{for almost all } \omega \in \Omega.$$

Such an  $\alpha$  can also be viewed as

$$\mu_n(\omega) \xrightarrow{\text{weak}} \alpha(\omega) \quad \text{for almost all } \omega \in \Omega,$$

where  $\mu_n = \frac{1}{n} \sum_{i=1}^n \delta_{X_i}$  is the empirical measure. Further,  $\alpha$  grants the usual representation

$$P(X \in B) = \int \alpha(\omega)^\infty(B) P(d\omega) \quad \text{for every Borel set } B \subset S^\infty$$

where  $\alpha(\omega)^\infty = \alpha(\omega) \times \alpha(\omega) \times \dots$

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Let  $\lambda$  be a  $\sigma$ -finite measure on  $\mathcal{B}$ . Our *first problem* is to give conditions for

$$(1) \quad \alpha(\omega) \ll \lambda \quad \text{for almost all } \omega \in \Omega.$$

The conditions should concern the finite dimensional distributions  $\mathcal{L}(X_1, \dots, X_n)$ ,  $n \geq 1$ , only.

While investigating (1), one meets another problem, of possible independent interest. Let  $\|\cdot\|$  denote total variation norm on  $(S, \mathcal{B})$ . Our *second problem* is to give conditions for

$$\|\alpha_n - \alpha\| \xrightarrow{a.s.} 0.$$

## 2. MOTIVATIONS

Again, let  $X = (X_1, X_2, \dots)$  be exchangeable.

Reasonable conditions for (1) look of theoretical interest. They are of practical interest as well thanks to Bayesian nonparametrics. In this framework, the starting point is a prior  $\pi$  on  $\mathbb{S}$ . Since  $\pi = P \circ \alpha^{-1}$ , condition (1) is equivalent to

$$\pi\{\nu \in \mathbb{S} : \nu \ll \lambda\} = 1.$$

This is a basic information for the subsequent statistical analysis. Roughly speaking, it means that the "underlying statistical model" consists of absolutely continuous laws.

Notwithstanding the significance of (1), however, there is a growing literature which gets around the first problem of this paper. Indeed, in a plenty of Bayesian nonparametric problems, condition (1) is just a crude *assumption* and the prior  $\pi$  is directly assessed on a set of densities (with respect to  $\lambda$ ). See e.g. [11] and references therein. Instead, it seems reasonable to get (1) as a consequence of explicit assumptions on the finite dimensional distributions  $\mathcal{L}(X_1, \dots, X_n)$ ,  $n \geq 1$ . From a foundational point of view, in fact, only assumptions on observable facts make sense. This attitude is strongly supported by de Finetti, among others. When dealing with the sequence  $X$ , the observable facts are events of the type  $\{(X_1, \dots, X_n) \in B\}$  for some  $n \geq 1$  and  $B \in \mathcal{B}^n$ . This is why, in this paper, the conditions for (1) are requested to concern  $\mathcal{L}(X_1, \dots, X_n)$ ,  $n \geq 1$ , only.

Some references related to the above remarks are [3] and [5]-[10]. In particular, in [8]-[9], Diaconis and Freedman have an exchangeable sequence of indicators and give conditions for the mixing measure (i.e., the prior  $\pi$ ) to be absolutely continuous with respect to Lebesgue measure. The present paper is much in the spirit of [8]-[9]. The main difference is that we give conditions for the mixands  $\{\alpha(\omega) : \omega \in \Omega\}$ , and not for the mixing measure  $\pi$ , to be absolutely continuous.

Next, a necessary condition for (1) is

$$(2) \quad \mathcal{L}(X_1, \dots, X_n) \ll \lambda^n \quad \text{for all } n \geq 1,$$

where  $\lambda^n = \lambda \times \dots \times \lambda$ . Condition (2) clearly involves the finite dimensional distributions only. Thus, a (natural) question is whether (2) suffices for (1) as well.

The answer is yes provided  $\alpha$  can be approximated by the predictive measures  $\alpha_n$  in some stronger sense. In fact, condition (2) can be written as

$$\alpha_n(\omega) \ll \lambda \quad \text{for all } n \geq 0 \text{ and almost all } \omega \in \Omega.$$

Hence, if (2) holds and  $\|\alpha_n - \alpha\| \xrightarrow{a.s.} 0$ , the set

$$A = \{\|\alpha_n - \alpha\| \rightarrow 0\} \cap \{\alpha_n \ll \lambda \text{ for all } n \geq 0\}$$

has probability 1. And, for each  $\omega \in A$ , one obtains

$$\alpha(\omega)(B) = \lim_n \alpha_n(\omega)(B) = 0 \quad \text{whenever } B \in \mathcal{B} \text{ and } \lambda(B) = 0.$$

Therefore, (1) follows from (2) and  $\|\alpha_n - \alpha\| \xrightarrow{a.s.} 0$ . In addition, a martingale argument implies the converse implication, that is

$$\alpha \ll \lambda \text{ a.s.} \iff \|\alpha_n - \alpha\| \xrightarrow{a.s.} 0 \text{ and } \mathcal{L}(X_1, \dots, X_n) \ll \lambda^n \text{ for all } n;$$

see Theorem 1. Thus, our first problem turns into the second one.

The question of whether  $\|\alpha_n - \alpha\| \xrightarrow{a.s.} 0$  is of independent interest. Among other things, it is connected to Bayesian consistency. Surprisingly, however, this question seems not answered so far. To the best of our knowledge,  $\|\alpha_n - \alpha\| \xrightarrow{a.s.} 0$  in every example known so far. And in fact, for some time, we conjectured that  $\|\alpha_n - \alpha\| \xrightarrow{a.s.} 0$  under condition (2). But this is not true. As shown in Example 5, when  $S = \mathbb{R}$  and  $\lambda =$  Lebesgue measure, it may be that  $\mathcal{L}(X_1, \dots, X_n) \ll \lambda^n$  for all  $n$  and yet  $\alpha$  is singular continuous a.s.. Indeed, the (topological) support of  $\alpha(\omega)$  has Hausdorff dimension 0 for almost all  $\omega \in \Omega$ .

Thus, (2) does not suffice for (1). To get (1), in addition to (2), one needs some growth conditions on the conditional densities. We refer to forthcoming Theorem 4 for such conditions. Here, we mention a result on the second problem. Actually, for  $\|\alpha_n - \alpha\| \xrightarrow{a.s.} 0$ , it suffices that

$$P\{\omega : \alpha_c(\omega) \ll \lambda\} = 1$$

where  $\alpha_c(\omega)$  denotes the continuous part of  $\alpha(\omega)$ ; see Theorem 2.

Finally, most results mentioned above do not need exchangeability of  $X$ , but the weaker assumption

$$(X_1, \dots, X_n, X_{n+2}) \sim (X_1, \dots, X_n, X_{n+1}) \quad \text{for all } n \geq 0.$$

Those sequences  $X$  satisfying the above condition, investigated in [2], are called *conditionally identically distributed* (c.i.d.).

### 3. MIXTURES OF I.I.D. ABSOLUTELY CONTINUOUS SEQUENCES

In this section,  $\mathcal{G}_0 = \{\emptyset, \Omega\}$ ,  $\mathcal{G}_n = \sigma(X_1, \dots, X_n)$  for  $n \geq 1$  and  $\mathcal{G}_\infty = \sigma(\cup_n \mathcal{G}_n)$ . If  $\mu$  is a random probability measure on  $S$ , we write  $\mu(B)$  to denote the real random variable  $\mu(\cdot)(B)$ ,  $B \in \mathcal{B}$ . Similarly, if  $h : S \rightarrow \mathbb{R}$  is a Borel function, integrable with respect to  $\mu(\omega)$  for almost all  $\omega \in \Omega$ , we write  $\mu(h)$  to denote  $\int h(x) \mu(\cdot)(dx)$ .

**3.1. Preliminaries.** Let  $X = (X_1, X_2, \dots)$  be c.i.d., as defined in Section 2. Since  $X$  needs not be exchangeable, the representation  $P(X \in \cdot) = \int \alpha(\omega)^\infty(\cdot) P(d\omega)$  can fail for any  $\alpha$ . However, there is a random probability measure  $\alpha$  on  $S$  such that

$$(3) \quad \sigma(\alpha) \subset \mathcal{G}_\infty \quad \text{and} \quad \alpha_n(B) = E\{\alpha(B) \mid \mathcal{G}_n\} \quad \text{a.s.}$$

for all  $B \in \mathcal{B}$ . In particular,  $\alpha_n \xrightarrow{weak} \alpha$  a.s.. Also, letting

$$\mu_n = \frac{1}{n} \sum_{i=1}^n \delta_{X_i}$$

be the empirical measure, one obtains  $\mu_n \xrightarrow{weak} \alpha$  a.s.. Such an  $\alpha$  is of interest for one more reason. There is an exchangeable sequence  $Y = (Y_1, Y_2, \dots)$  of  $S$ -valued random variables on  $(\Omega, \mathcal{A}, P)$  such that

$$(X_n, X_{n+1}, \dots) \xrightarrow{d} Y \quad \text{and} \quad P(Y \in \cdot) = \int \alpha(\omega)^\infty(\cdot) P(d\omega).$$

See [2] for details.

We next recall some known facts about vector-valued martingales; see [14]. Let  $(\mathcal{Z}, \|\cdot\|_*)$  be a separable Banach space. Also, let  $\mathcal{F} = (\mathcal{F}_n)$  be a filtration and  $(Z_n)$  a sequence of  $\mathcal{Z}$ -valued random variables on  $(\Omega, \mathcal{A}, P)$  such that  $E\|Z_n\|_* < \infty$  for all  $n$ . Then,  $(Z_n)$  is an  $\mathcal{F}$ -martingale in case  $(\phi(Z_n))$  is an  $\mathcal{F}$ -martingale for each linear continuous functional  $\phi : \mathcal{Z} \rightarrow \mathbb{R}$ . If  $(Z_n)$  is an  $\mathcal{F}$ -martingale,  $(\|Z_n\|_*)$  is a real-valued  $\mathcal{F}$ -submartingale. So, Doob's maximal inequality yields

$$E\left\{\sup_n \|Z_n\|_*^p\right\} \leq \left(\frac{p}{p-1}\right)^p \sup_n E\{\|Z_n\|_*^p\} \quad \text{for all } p > 1.$$

The following martingale convergence theorem is available as well. Let  $Z : \Omega \rightarrow \mathcal{Z}$  be  $\mathcal{F}_\infty$ -measurable and such that  $E\|Z\|_* < \infty$ , where  $\mathcal{F}_\infty = \sigma(\cup_n \mathcal{F}_n)$ . Then,  $Z_n \xrightarrow{a.s.} Z$  provided  $\phi(Z_n) = E\{\phi(Z) | \mathcal{F}_n\}$  a.s. for all  $n$  and all linear continuous functionals  $\phi : \mathcal{Z} \rightarrow \mathbb{R}$ .

**3.2. Results.** In the sequel,  $\lambda$  is a  $\sigma$ -finite measure on  $\mathcal{B}$ . When  $S = \mathbb{R}$ , it may be natural to think of  $\lambda$  as the Lebesgue measure, but this is only a particular case. Indeed,  $\lambda$  could be singular continuous or concentrated on any Borel subset. In addition,  $X$  is c.i.d. (in particular, exchangeable) and  $\alpha$  is a random probability measure on  $S$  such that  $\alpha_n \xrightarrow{weak} \alpha$  a.s.. Equivalently,  $\alpha$  can be obtained as  $\mu_n \xrightarrow{weak} \alpha$  a.s.. It can (and will) be assumed  $\sigma(\alpha) \subset \mathcal{G}_\infty$ .

**Theorem 1.** *Suppose  $X = (X_1, X_2, \dots)$  is c.i.d.. Then,  $\alpha \ll \lambda$  a.s. if and only if  $\|\alpha_n - \alpha\| \xrightarrow{a.s.} 0$  and  $\mathcal{L}(X_1, \dots, X_n) \ll \lambda^n$  for all  $n$ .*

*Proof.* The "if" part can be proved exactly as in Section 2. Conversely, suppose  $\alpha \ll \lambda$  a.s.. It can be assumed  $\alpha(\omega) \ll \lambda$  for all  $\omega \in \Omega$ . We let  $L_p = L_p(S, \mathcal{B}, \lambda)$  for each  $1 \leq p \leq \infty$ .

Let  $f : \Omega \times S \rightarrow [0, \infty)$  be such that  $\alpha(\omega)(dx) = f(\omega, x) \lambda(dx)$  for all  $\omega \in \Omega$ . Since  $\mathcal{B}$  is countably generated,  $f$  can be taken to be  $\mathcal{A} \otimes \mathcal{B}$ -measurable (see [4], V.5.58, page 52) so that

$$1 = \int 1 dP = \int \int f(\omega, x) \lambda(dx) P(d\omega) = \int E\{f(\cdot, x)\} \lambda(dx).$$

Thus, given  $n \geq 0$ ,  $E\{f(\cdot, x) | \mathcal{G}_n\}$  is well defined for  $\lambda$ -almost all  $x \in S$ . Since  $X$  is c.i.d., condition (3) also implies

$$\begin{aligned} \int_B E\{f(\cdot, x) | \mathcal{G}_n\} \lambda(dx) &= E\left\{\int_B f(\cdot, x) \lambda(dx) | \mathcal{G}_n\right\} \\ &= E\{\alpha(B) | \mathcal{G}_n\} = \alpha_n(B) \quad \text{a.s. for fixed } B \in \mathcal{B}. \end{aligned}$$

Since  $\mathcal{B}$  is countably generated, the previous equality yields

$$\alpha_n(\omega)(dx) = E\{f(\cdot, x) | \mathcal{G}_n\}(\omega) \lambda(dx) \quad \text{for almost all } \omega \in \Omega.$$

This proves that  $\mathcal{L}(X_1, \dots, X_n) \ll \lambda^n$  for all  $n$ . In particular, up to modifying  $\alpha_n$  on a  $P$ -null set, it can be assumed  $\alpha_n(\omega)(dx) = f_n(\omega, x) \lambda(dx)$  for all  $n \geq 0$ , all  $\omega \in \Omega$ , and suitable functions  $f_n : \Omega \times S \rightarrow [0, \infty)$ .

Regard  $f, f_n : \Omega \rightarrow L_1$  as  $L_1$ -valued random variables. Then,  $f : \Omega \rightarrow L_1$  is  $\mathcal{G}_\infty$ -measurable for  $\int h(x) f(\cdot, x) \lambda(dx) = \alpha(h)$  is  $\mathcal{G}_\infty$ -measurable for all  $h \in L_\infty$ . Clearly,  $\|f(\omega, \cdot)\|_{L_1} = \|f_n(\omega, \cdot)\|_{L_1} = 1$  for all  $n$  and  $\omega$ . Finally,  $X$  c.i.d. implies

$$\begin{aligned} E\left\{\int h(x) f(\cdot, x) \lambda(dx) \mid \mathcal{G}_n\right\} &= E\{\alpha(h) \mid \mathcal{G}_n\} = \alpha_n(h) \\ &= \int h(x) f_n(\cdot, x) \lambda(dx) \quad \text{a.s. for all } h \in L_\infty. \end{aligned}$$

By the martingale convergence theorem (see Subsection 3.1)  $f_n \xrightarrow{\text{a.s.}} f$  in the space  $L_1$ , that is

$$\|\alpha_n(\omega) - \alpha(\omega)\| = \frac{1}{2} \int |f_n(\omega, x) - f(\omega, x)| \lambda(dx) \longrightarrow 0 \quad \text{for almost all } \omega \in \Omega.$$

□

In the exchangeable case, the argument of the previous proof yields a little bit more. Indeed, if  $X$  is exchangeable and  $\alpha \ll \lambda$  a.s., then

$$\sup_{B \in \mathcal{B}^k} \left| P\{(X_{n+1}, \dots, X_{n+k}) \in B \mid \mathcal{G}_n\} - \alpha^k(B) \right| \xrightarrow{\text{a.s.}} 0,$$

where  $k \geq 1$  is any integer and  $\alpha^k = \alpha \times \dots \times \alpha$ .

The next result deals with the second problem of Section 1. For each  $\nu \in \mathbb{S}$ , let  $\nu_c$  and  $\nu_d$  denote the continuous and discrete parts of  $\nu$ , that is,  $\nu_d(B) = \sum_{x \in B} \nu\{x\}$  for all  $B \in \mathcal{B}$  and  $\nu_c = \nu - \nu_d$ .

**Theorem 2.** *Suppose  $X = (X_1, X_2, \dots)$  is c.i.d. and  $P\{\omega : \alpha_c(\omega) \ll \lambda\} = 1$ . Then,  $\|\alpha_n - \alpha\| \xrightarrow{\text{a.s.}} 0$  if and only if*

$$(4) \quad \begin{aligned} &\text{there is a set } A_0 \in \mathcal{A} \text{ such that } P(A_0) = 1 \text{ and} \\ &\alpha_n(\omega)\{x\} \longrightarrow \alpha(\omega)\{x\} \text{ for all } x \in S \text{ and } \omega \in A_0. \end{aligned}$$

(Recall that  $\mathcal{A}$  denotes the basic  $\sigma$ -field on  $\Omega$ ). Moreover, condition (4) is automatically true if  $X$  is exchangeable, so that  $\|\alpha_n - \alpha\| \xrightarrow{\text{a.s.}} 0$  provided  $X$  is exchangeable and  $\alpha_c \ll \lambda$  a.s..

*Proof.* The "only if" part is trivial. Suppose condition (4) holds. For each  $n \geq 0$ , take functions  $\beta_n$  and  $\gamma_n$  on  $\Omega$  such that  $\beta_n(\omega)$  and  $\gamma_n(\omega)$  are measures on  $\mathcal{B}$  for all  $\omega \in \Omega$  and

$$\beta_n(B) = E\{\alpha_c(B) \mid \mathcal{G}_n\}, \quad \gamma_n(B) = E\{\alpha_d(B) \mid \mathcal{G}_n\}, \quad \text{a.s.},$$

for all  $B \in \mathcal{B}$ . Since  $X$  is c.i.d., condition (3) yields  $\alpha_n = \beta_n + \gamma_n$  a.s..

We first prove  $\|\beta_n - \alpha_c\| \xrightarrow{\text{a.s.}} 0$ . It can be assumed  $\alpha_c(\omega) \ll \lambda$  for all  $\omega \in \Omega$ , so that  $\alpha_c(\omega)(dx) = f(\omega, x) \lambda(dx)$  for all  $\omega \in \Omega$  and some function  $f : \Omega \times S \rightarrow [0, \infty)$ . For fixed  $B \in \mathcal{B}$ , arguing as in the proof of Theorem 1, one has

$$\beta_n(B) = E\left\{\int_B f(\cdot, x) \lambda(dx) \mid \mathcal{G}_n\right\} = \int_B E(f(\cdot, x) \mid \mathcal{G}_n) \lambda(dx) \quad \text{a.s.}$$

By standard arguments, it follows that  $\beta_n \ll \lambda$  a.s.. Again, it can be assumed  $\beta_n(\omega)(dx) = f_n(\omega, x) \lambda(dx)$  for all  $\omega \in \Omega$  and some function  $f_n : \Omega \times S \rightarrow [0, \infty)$ .

Define  $L_1 = L_1(S, \mathcal{B}, \lambda)$  and regard  $f_n, f : \Omega \rightarrow L_1$  as  $L_1$ -valued random variables. By the same martingale argument used for Theorem 1, one obtains  $f_n \xrightarrow{a.s.} f$  in the space  $L_1$ . That is,  $\|\beta_n - \alpha_c\| \xrightarrow{a.s.} 0$ .

We next prove  $\|\gamma_n - \alpha_d\| \xrightarrow{a.s.} 0$ . Take  $A_0$  as in condition (4) and define

$$A_1 = \left\{ \lim_n \|f_n - f\|_{L_1} = 0 \text{ and } \alpha_n = \beta_n + \gamma_n \text{ for all } n \geq 0 \right\}.$$

Then,  $P(A_0 \cap A_1) = 1$  and

$$\begin{aligned} \alpha_d(\omega)\{x\} &= \alpha(\omega)\{x\} - \alpha_c(\omega)\{x\} = \alpha(\omega)\{x\} - f(\omega, x) \lambda\{x\} \\ &= \lim_n (\alpha_n(\omega)\{x\} - f_n(\omega, x) \lambda\{x\}) = \lim_n (\alpha_n(\omega)\{x\} - \beta_n(\omega)\{x\}) = \lim_n \gamma_n(\omega)\{x\} \end{aligned}$$

for all  $\omega \in A_0 \cap A_1$  and  $x \in S$ . Define also

$$A = A_0 \cap A_1 \cap \{\gamma_n(S) \rightarrow \alpha_d(S)\}.$$

Since  $\gamma_n(S) = 1 - \beta_n(S) \xrightarrow{a.s.} 1 - \alpha_c(S) = \alpha_d(S)$ , then  $P(A) = 1$ . Fix  $\omega \in A$  and let  $D_\omega = \{x \in S : \alpha(\omega)\{x\} > 0\}$ . Then,

$$\alpha_d(\omega)(D_\omega) \leq \liminf_n \gamma_n(\omega)(D_\omega)$$

since  $D_\omega$  is countable and  $\alpha_d(\omega)\{x\} = \lim_n \gamma_n(\omega)\{x\}$  for all  $x \in D_\omega$ . Further,

$$\limsup_n \gamma_n(\omega)(D_\omega) \leq \limsup_n \gamma_n(\omega)(S) = \alpha_d(\omega)(S) = \alpha_d(\omega)(D_\omega).$$

Therefore,  $\lim_n \|\gamma_n(\omega) - \alpha_d(\omega)\| = 0$  is an immediate consequence of

$$\begin{aligned} \gamma_n(\omega)\{x\} &\rightarrow \alpha_d(\omega)\{x\} \quad \text{for each } x \in D_\omega, \\ \alpha_d(\omega)(D_\omega) &= \lim_n \gamma_n(\omega)(D_\omega), \quad \alpha_d(\omega)(D_\omega^c) = \lim_n \gamma_n(\omega)(D_\omega^c) = 0. \end{aligned}$$

Finally, suppose  $X$  is exchangeable. We have to prove condition (4). If  $S$  is countable, condition (4) is trivial for  $\alpha_n(B) \xrightarrow{a.s.} \alpha(B)$  for fixed  $B \in \mathcal{B}$ . If  $S = \mathbb{R}$ , Glivenko-Cantelli theorem yields  $\sup_x |\mu_n(I_x) - \alpha(I_x)| \xrightarrow{a.s.} 0$ , where  $I_x = (-\infty, x]$  and  $\mu_n = \frac{1}{n} \sum_{i=1}^n \delta_{X_i}$  is the empirical measure. Hence, (4) follows from

$$\sup_x |\alpha_n(I_x) - \mu_n(I_x)| \xrightarrow{a.s.} 0;$$

see Corollary 3.2 of [1]. If  $S$  is any uncountable Polish space, take a Borel isomorphism  $\psi : S \rightarrow \mathbb{R}$ . (Thus,  $\psi$  is bijective with  $\psi$  and  $\psi^{-1}$  Borel measurable). Then,  $(\psi(X_n))$  is an exchangeable sequence of real random variables and condition (4) is a straightforward consequence of

$$P\{\psi(X_{n+1}) \in B \mid \psi(X_1), \dots, \psi(X_n)\} = P\{\psi(X_{n+1}) \in B \mid \mathcal{G}_n\} = \alpha_n(\psi^{-1}B) \quad \text{a.s.}$$

for each Borel set  $B \subset \mathbb{R}$ . This concludes the proof.  $\square$

When  $X$  is c.i.d. (but not exchangeable)  $\|\alpha_n - \alpha\| \xrightarrow{a.s.} 0$  needs not be true even if  $\alpha_c \ll \lambda$  a.s..

**Example 3.** Let  $(Z_n)$  and  $(U_n)$  be independent sequences of independent real random variables such that  $Z_n \sim \mathcal{N}(0, b_n - b_{n-1})$  and  $U_n \sim \mathcal{N}(0, 1 - b_n)$ , where  $0 = b_0 < b_1 < b_2 < \dots < 1$  and  $\sum_n (1 - b_n) < \infty$ . As shown in Example 1.2 of [2],

$$X_n = \sum_{i=1}^n Z_i + U_n$$

is c.i.d. and  $X_n \xrightarrow{a.s.} V$  for some real random variable  $V$ . Since  $\mu_n \xrightarrow{weak} \delta_V$  a.s., then  $\alpha = \delta_V$  and  $\alpha_c \ll \lambda$  a.s. (in fact,  $\alpha_c = 0$  a.s.). However, condition (4) fails. In fact,  $\mathcal{L}(X_1, \dots, X_n) \ll \lambda^n$  for all  $n$ , where  $\lambda$  is Lebesgue measure. Hence,  $\alpha_n(\omega)\{V(\omega)\} = 0$  while  $\alpha(\omega)\{V(\omega)\} = 1$  for all  $n$  and almost all  $\omega \in \Omega$ .

We now turn to the first problem of Section 1. Recall that condition (2) amounts to  $\alpha_n \ll \lambda$  a.s. for all  $n \geq 0$ . Therefore, up to modifying  $\alpha_n$  on a  $P$ -null set, under condition (2) one can write

$$\alpha_n(\omega)(dx) = f_n(\omega, x) \lambda(dx)$$

for each  $\omega \in \Omega$ , each  $n \geq 0$ , and some function  $f_n : \Omega \times S \rightarrow [0, \infty)$ . We also let

$$\begin{aligned} \mathcal{K} &= \{K : K \text{ compact subset of } S \text{ and } \lambda(K) < \infty\} \\ \text{and } \lambda_B(\cdot) &= \lambda(\cdot \cap B) \quad \text{for all } B \in \mathcal{B}. \end{aligned}$$

**Theorem 4.** *Suppose  $X = (X_1, X_2, \dots)$  is c.i.d. and  $\mathcal{L}(X_1, \dots, X_n) \ll \lambda^n$  for all  $n$ . Then,  $\alpha \ll \lambda$  a.s. if and only if, for each  $K \in \mathcal{K}$ ,*

$$(5) \quad \begin{aligned} &\text{the sequence } (f_n(\omega, \cdot) : n \geq 1) \text{ is uniformly integrable,} \\ &\text{in the space } (S, \mathcal{B}, \lambda_K), \text{ for almost all } \omega \in \Omega. \end{aligned}$$

*In particular,  $\alpha \ll \lambda$  a.s. provided, for each  $K \in \mathcal{K}$ , there is  $p > 1$  such that*

$$(6) \quad \sup_n \int_K f_n(\omega, x)^p \lambda(dx) < \infty \quad \text{for almost all } \omega \in \Omega.$$

*Moreover, for condition (6) to be true, it suffices that*

$$\sup_n E \left\{ \int_K f_n^p d\lambda \right\} < \infty.$$

*Proof.* If  $\alpha \ll \lambda$  a.s., Theorem 1 yields  $\|\alpha_n - \alpha\| \xrightarrow{a.s.} 0$ . Thus,  $f_n(\omega, \cdot)$  converges in  $L_1(S, \mathcal{B}, \lambda)$ , for almost all  $\omega \in \Omega$ , and this implies condition (5). Conversely, we now prove that  $\alpha \ll \lambda$  a.s. under condition (5).

Fix a nondecreasing sequence  $B_1 \subset B_2 \subset \dots$  such that  $B_n \in \mathcal{B}$ ,  $\lambda(B_n) < \infty$ , and  $\cup_n B_n = S$ . Since  $\lambda(B_1) < \infty$  and  $S$  is Polish, there is  $K_1 \in \mathcal{K}$  satisfying  $K_1 \subset B_1$  and  $\lambda(B_1 \cap K_1^c) < 1$ . By induction, for each  $n \geq 2$ , there is  $K_n \in \mathcal{K}$  such that  $K_{n-1} \subset K_n \subset B_n$  and  $\lambda(B_n \cap K_n^c) < 1/n$ . Since  $X$  is c.i.d., condition (3) implies

$$\alpha(K_m) = \lim_n E \{ \alpha(K_m) \mid \mathcal{G}_n \} = \lim_n \alpha_n(K_m) \quad \text{a.s. for all } m \geq 1.$$

Define  $H = \cup_m K_m$  and  $A_H = \{\alpha(H) = 1\}$ . If  $\omega \in A_H$ , then

$$\alpha(\omega)(B) = \alpha(\omega)(B \cap H) = \sup_m \alpha(\omega)(B \cap K_m) \quad \text{for all } B \in \mathcal{B}.$$

Moreover,  $P(A_H) = 1$ . In fact,  $\lambda(H^c) = 0$  and  $\alpha_n \ll \lambda$  a.s. for all  $n$ , so that

$$\alpha(H) = \lim_n E \{ \alpha(H) \mid \mathcal{G}_n \} = \lim_n \alpha_n(H) = 1 \quad \text{a.s..}$$

Thus, to prove  $\alpha \ll \lambda$  a.s., it suffices to see that  $\alpha(\cdot \cap K_m) \ll \lambda$  a.s. for all  $m$ .

Suppose (5) holds. Fix  $m \geq 1$ , define  $K = K_m$ , and take a set  $A \in \mathcal{A}$  such that  $P(A) = 1$  and, for each  $\omega \in A$ ,

$$\begin{aligned} \alpha(\omega)(K) &= \lim_n \alpha_n(\omega)(K), \quad \alpha_n(\omega) \xrightarrow{weak} \alpha(\omega), \\ (f_n(\omega, \cdot) : n \geq 1) &\text{ is uniformly integrable in } (S, \mathcal{B}, \lambda_K). \end{aligned}$$

Let  $\omega \in A$ . Since  $\lambda_K(S) = \lambda(K) < \infty$  and  $(f_n(\omega, \cdot) : n \geq 1)$  is uniformly integrable under  $\lambda_K$ , there is a subsequence  $(n_j)$  and a function  $\psi_\omega \in L_1(S, \mathcal{B}, \lambda_K)$  such that  $f_{n_j}(\omega, \cdot) \rightarrow \psi_\omega$  in the weak-topology of  $L_1(S, \mathcal{B}, \lambda_K)$ . This means that

$$\int_{B \cap K} \psi_\omega(x) \lambda(dx) = \lim_j \int_{B \cap K} f_{n_j}(\omega, x) \lambda(dx) = \lim_j \alpha_{n_j}(\omega)(B \cap K) \quad \text{for all } B \in \mathcal{B}.$$

Therefore,

$$\begin{aligned} \int_K \psi_\omega(x) \lambda(dx) &= \lim_j \alpha_{n_j}(\omega)(K) = \alpha(\omega)(K) \quad \text{and} \\ \int_{F \cap K} \psi_\omega(x) \lambda(dx) &= \lim_j \alpha_{n_j}(\omega)(F \cap K) \leq \alpha(\omega)(F \cap K) \quad \text{for each closed } F \subset S. \end{aligned}$$

By standard arguments, the previous two relations yield

$$\alpha(\omega)(B \cap K) = \int_{B \cap K} \psi_\omega(x) \lambda(dx) \quad \text{for all } B \in \mathcal{B}.$$

Thus,  $\alpha(\omega)(\cdot \cap K) \ll \lambda$ . This proves that condition (5) implies  $\alpha \ll \lambda$  a.s..

Next, since  $p > 1$ , it is obvious that (6)  $\implies$  (5). Hence, it remains only to see that condition (6) follows from  $\sup_n E \left\{ \int_K f_n^p d\lambda \right\} < \infty$ .

Fix  $B \in \mathcal{B}$ ,  $p > 1$ , and suppose  $\sup_n E \left\{ \int_B f_n^p d\lambda \right\} < \infty$ . Let  $L_r = L_r(S, \mathcal{B}, \lambda_B)$  for all  $r$ . It can be assumed  $\int_B f_n(\omega, x)^p \lambda(dx) < \infty$  for all  $\omega \in \Omega$  and  $n \geq 1$ . Thus, each  $f_n : \Omega \rightarrow L_p$  can be seen as an  $L_p$ -valued random variable such that

$$E \{ \|f_n\|_{L_p} \} = E \left\{ \left( \int_B f_n^p d\lambda \right)^{1/p} \right\} \leq \left( E \left\{ \int_B f_n^p d\lambda \right\} \right)^{1/p} < \infty.$$

Further,  $\int f_n(\cdot, x) h(x) \lambda_B(dx) = \alpha_n(I_B h)$  is  $\mathcal{G}_n$ -measurable for all  $h \in L_q$ , where  $q = p/(p-1)$ . Since  $X$  is c.i.d., condition (3) also implies

$$\begin{aligned} E \left\{ \int f_{n+1}(\cdot, x) h(x) \lambda_B(dx) \mid \mathcal{G}_n \right\} &= E \{ \alpha_{n+1}(I_B h) \mid \mathcal{G}_n \} \\ &= E \left\{ E(\alpha(I_B h) \mid \mathcal{G}_{n+1}) \mid \mathcal{G}_n \right\} \\ &= E \{ \alpha(I_B h) \mid \mathcal{G}_n \} = \alpha_n(I_B h) \\ &= \int f_n(\cdot, x) h(x) \lambda_B(dx) \quad \text{a.s. for all } h \in L_q. \end{aligned}$$

Hence,  $(f_n)$  is a  $(\mathcal{G}_n)$ -martingale. By Doob's maximal inequality,

$$\begin{aligned} E \left\{ \sup_n \int_B f_n^p d\lambda \right\} &= E \left\{ \sup_n \|f_n\|_{L_p}^p \right\} \\ &\leq q^p \sup_n E \left\{ \|f_n\|_{L_p}^p \right\} = q^p \sup_n E \left\{ \int_B f_n^p d\lambda \right\} < \infty. \end{aligned}$$

In particular,  $\sup_n \int_B f_n^p d\lambda < \infty$  a.s., and this concludes the proof.  $\square$

Some remarks on Theorem 4 are in order.

First, for  $S = [0, 1]$  and a particular class of exchangeable sequences, results similar to Theorem 4 are in [12] and [13].



Second,

$$f_n(\omega, \cdot) = \frac{g_{n+1}(X_1(\omega), \dots, X_n(\omega), \cdot)}{g_n(X_1(\omega), \dots, X_n(\omega))} \quad \text{for almost all } \omega \in \Omega,$$

where each  $g_n : S^n \rightarrow [0, \infty)$  is a density of  $\mathcal{L}(X_1, \dots, X_n)$  with respect to  $\lambda^n$ . Thus, more concretely, one obtains

$$\int_K f_n^p d\lambda = \frac{\int_K g_{n+1}(X_1, \dots, X_n, x)^p \lambda(dx)}{g_n(X_1, \dots, X_n)^p} \quad \text{a.s.}$$

Third, suppose  $X$  exchangeable and fix *any* random probability measure  $\gamma$  on  $S$  such that  $P(X \in \cdot) = \int \gamma(\omega)^\infty(\cdot) P(d\omega)$ . Then,  $\gamma \ll \lambda$  a.s. under the assumptions of Theorem 4. In fact,  $\alpha$  and  $\gamma$  have the same probability distribution, when regarded as  $\mathbb{S}$ -valued random variables.

A last (and important) remark deals with condition (2). Indeed, even if  $X$  is exchangeable, condition (2) is not enough for  $\alpha \ll \lambda$  a.s.. We close the paper showing this fact.

**Example 5.** Let  $S = \mathbb{R}$  and  $\lambda =$  Lebesgue measure. All random variables are defined on the probability space  $(\Omega, \mathcal{A}, P)$ . We now exhibit an exchangeable sequence  $X$  such that  $\mathcal{L}(X_1, \dots, X_n) \ll \lambda^n$  for all  $n \geq 1$  and yet  $P(\alpha \ll \lambda) = 0$ . In fact, the support of  $\alpha(\omega)$  has Hausdorff dimension 0 for almost all  $\omega \in \Omega$ .

Two known facts are to be recalled. First, if  $T$  and  $Z$  are independent  $\mathbb{R}^n$ -valued random variables, then

$$P(T + Z \in B) = \int P(T + z \in B) P_Z(dz)$$

where  $B \in \mathcal{B}^n$  and  $P_Z$  is the distribution of  $Z$ . Hence,  $\mathcal{L}(T + Z) \ll \lambda^n$  provided  $\mathcal{L}(T) \ll \lambda^n$ . The second fact is

**Theorem 6. (Pratsiovytyi and Feshchenko).** *Let  $Z_1, Z_2, \dots$  be i.i.d. random variables with  $P(Z_1 = 0) = P(Z_1 = 1) = 1/2$  and  $b_1 > b_2 > \dots > 0$  real numbers such that  $\sum_m b_m < \infty$ . Then, the support of  $\mathcal{L}(\sum_m b_m Z_m)$  has Hausdorff dimension 0 whenever  $\lim_m (\sum_{j>m} b_j)^{-1} b_m = \infty$ .*

Theorem 6 is a consequence of Theorem 8 of [15] (which is actually much more general).

Next, let  $U_m$  and  $Y_{m,n}$  be independent real random variables such that:

- $U_m$  is uniformly distributed on  $(\frac{1}{m+1}, \frac{1}{m})$  for each  $m \geq 1$ ;
- $P(Y_{m,n} = 0) = P(Y_{m,n} = 1) = \frac{1}{2}$  for all  $m, n \geq 1$ .

Define  $V_m = U_m^m$  and

$$X_n = \sum_{m=1}^{\infty} U_m^m Y_{m,n} = \sum_{m=1}^{\infty} V_m Y_{m,n}.$$

Then,  $X = (X_1, X_2, \dots)$  is conditionally i.i.d. given  $\mathcal{V} = \sigma(V_1, V_2, \dots)$ . Precisely, for  $\omega \in \Omega$  and  $B \in \mathcal{B}$ , define

$$\alpha(\omega)(B) = P\left\{u \in \Omega : \sum_m V_m(\omega) Y_{m,1}(u) \in B\right\}.$$

Then,  $\alpha(B)$  is a version of  $P(X_1 \in B \mid \mathcal{V})$  and  $P(X \in \cdot) = \int \alpha(\omega)^\infty(\cdot) P(d\omega)$ . In particular,  $X$  is exchangeable. Moreover,  $\mu_n \xrightarrow{weak} \alpha$  a.s. for

$$P(\mu_n \xrightarrow{weak} \alpha \mid \mathcal{V}) = 1 \quad \text{a.s.}$$

The (topological) support of  $\alpha(\omega)$  has Hausdorff dimension 0 for almost all  $\omega \in \Omega$ . Define in fact  $b_m = V_m(\omega)$  and  $Z_m = Y_{m,1}$ . By Theorem 6, it suffices to verify that

$$(7) \quad \lim_m \frac{V_m(\omega)}{\sum_{j>m} V_j(\omega)} = \infty \quad \text{for almost all } \omega \in \Omega.$$

And condition (7) follows immediately from

$$(j+1)^{-j} < V_j < j^{-j} \quad \text{and} \quad \sum_{j>m} V_j \leq \sum_{j>m} j^{-j} \leq \sum_{j>m} (m+1)^{-j} = \frac{(m+1)^{-m}}{m} \quad \text{a.s.}$$

We finally prove that  $\mathcal{L}(X_1, \dots, X_n) \ll \lambda^n$  for all  $n \geq 1$ . Given the array  $y = (y_{m,n} : m, n \geq 1)$ , with  $y_{m,n} \in \{0, 1\}$  for all  $m, n$ , define

$$X_{n,y} = \sum_m V_m y_{m,n}.$$

Fix  $n \geq 1$  and denote  $I_n$  the  $n \times n$  identity matrix. If  $y$  satisfies

$$(8) \quad \begin{pmatrix} y_{m+1,1} & \cdots & y_{m+1,n} \\ \cdots & \cdots & \cdots \\ y_{m+n,1} & \cdots & y_{m+n,n} \end{pmatrix} = I_n \quad \text{for some } m \geq 0,$$

then

$$(X_{1,y}, \dots, X_{n,y}) = (V_{m+1}, \dots, V_{m+n}) + (R_1, \dots, R_n) \\ \text{with } (R_1, \dots, R_n) \text{ independent of } (V_{m+1}, \dots, V_{m+n}).$$

In this case, since  $\mathcal{L}(V_{m+1}, \dots, V_{m+n}) \ll \lambda^n$ , then  $\mathcal{L}(X_{1,y}, \dots, X_{n,y}) \ll \lambda^n$ . Hence, letting  $Y = (Y_{m,n} : m, n \geq 1)$ , the conditional distribution of  $(X_1, \dots, X_n)$  given  $Y = y$  is absolutely continuous with respect to  $\lambda^n$  as far as  $y$  satisfies (8). To conclude the proof, it suffices noting that

$$P(Y = y \text{ for some } y \text{ satisfying (8)}) = 1.$$

#### REFERENCES

- [1] Berti P., Mattei A., Rigo P. (2002) Uniform convergence of empirical and predictive measures, *Atti Sem. Mat. Fis. Univ. Modena*, L, 465-477.
- [2] Berti P., Pratelli L., Rigo P. (2004) Limit theorems for a class of identically distributed random variables, *Ann. Probab.*, 32, 2029-2052.
- [3] Cifarelli D.M., Regazzini E. (1996) de Finetti's contribution to probability and statistics, *Statist. Sci.*, 11, 253-282.
- [4] Dellacherie C., Meyer P.A. (1982) *Probabilities and Potential B*, North-Holland.
- [5] Diaconis P., Freedman D. (1984) Partial exchangeability and sufficiency, *Proc. of the Indian Statistical Institute Golden Jubilee International Conference on Statistics: Applications and New Directions* (J. K. Ghosh and J. Roy, eds.) 205-236. Indian Statistical Institute, Calcutta.

- [6] Diaconis P., Freedman D. (1988) Conditional limit theorems for exponential families and finite versions of de Finetti's theorem, *J. Theoret. Probab.*, 1, 381-410.
- [7] Diaconis P., Freedman D. (1990) Cauchy's equation and de Finetti's theorem, *Scandinavian J. Statist.*, 17, 235-274.
- [8] Diaconis P., Freedman D. (2004) The Markov moment problem and de Finetti's theorem Part I, *Mathematische Zeitschrift*, 247(1), 183-199.
- [9] Diaconis P., Freedman D. (2004) The Markov moment problem and de Finetti's theorem Part II, *Mathematische Zeitschrift*, 247(1), 201-212.
- [10] Fortini S., Ladelli L., Regazzini E. (2000) Exchangeability, predictive distributions and parametric models, *Sankhya A*, 62, 86-109.
- [11] Ghosal S., van der Vaart A.W. (2012) *Fundamentals of nonparametric Bayesian inference*, Cambridge University Press, to appear.
- [12] Kraft C.H. (1964) A class of distribution function processes which have derivatives, *J. Appl. Probab.*, 1, 385-388.
- [13] Metivier M. (1971) Sur la construction de mesures aleatoires presque surement absolument continues par rapport a une mesure donnee, *Z. Wahrsch. verw. Geb.*, 20, 332-344.
- [14] Neveu J. (1975) *Discrete parameter martingales*, North-Holland.
- [15] Pratsiovytyi M.V., Feshchenko O.YU. (2007) Topological, metric and fractal properties of probability distributions on the set of incomplete sums of positive series, *Theory Stoch. Proc.*, 13, 205-224.

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